

ADDITIVE PERTURBATION

For all $n \geq 1$, we consider the following $n \times n$ random matrix:

$$W_n = \frac{1}{\sqrt{n}} X_n + \theta e_1 e_1^T,$$

where $\theta \geq 0$ and

- X_n is a symmetric $n \times n$ random matrix with i.i.d. entries that are centered and reduced;
- e_1 is the first vector of the canonical basis.

Let $\lambda_1(W_n) \geq \dots \geq \lambda_n(W_n)$ be the eigenvalues of W_n and $\phi_1(W_n), \dots, \phi_n(W_n)$ the associated eigenvectors.

MULTIPLICATIVE PERTURBATION

For all $n \geq 1$, we consider the following $n \times n$ random matrix:

$$S_n = \frac{1}{n} \Sigma_n^{1/2} X_n X_n^T \Sigma_n^{1/2},$$

where $\theta \geq 1$ and

- X_n is a $n \times \alpha n$ ($\alpha \geq 1$) random matrix with i.i.d. entries that are centered and reduced;
- $\Sigma_n = \text{Diag}(\theta, 1, \dots, 1)$ is of size $n \times n$.

Let $\lambda_1(S_n) \geq \dots \geq \lambda_n(S_n)$ be the eigenvalues of S_n and $\phi_1(S_n), \dots, \phi_n(S_n)$ the associated eigenvectors.

SPECTRAL MEASURES IN DIRECTION e_1

Heuristically, when $\theta \gg 1$, it should be close to an eigenvalue of W_n (resp. S_n) and its eigenvector should be close to e_1 . In order to reveal the influence of θ at a macroscopical level, one should therefore look at a statistic which *projects* the spectrum of W_n (resp. S_n) onto e_1 . This is precisely the definition of the spectral measures in direction e_1 :

$$\mu_{(W_n, e_1)} := \sum_{i=1}^n |\langle \phi_i(W_n), e_1 \rangle|^2 \delta_{\lambda_i(W_n)} \quad \text{and} \quad \mu_{(S_n, e_1)} := \sum_{i=1}^n |\langle \phi_i(S_n), e_1 \rangle|^2 \delta_{\lambda_i(S_n)}.$$

Their Stieltjes transforms are given by $\langle e_1, (W_n - z)^{-1} e_1 \rangle$ and $\langle e_1, (S_n - z)^{-1} e_1 \rangle$. These quantities have been carefully analyzed, the most recent results being the local laws obtained by Knowles and Yin in [1]. For fixed z , it implies the pointwise convergence of the Stieltjes transforms. Interestingly, the associated limiting probability measures are explicit, given respectively by:

$$\mu_{sc, \theta}(dx) = \frac{\sqrt{4 - x^2} \mathbf{1}_{|x| \leq 2} dx}{2\pi(\theta^2 - \theta x + 1)} + \mathbf{1}_{|\theta| > 1} \left(1 - \frac{1}{\theta^2}\right) \delta_{\theta + \frac{1}{\theta}}(dx) \quad \text{and} \quad \mu_{\alpha, \theta}(dx) = \frac{\theta \sqrt{(b-x)(x-a)} \mathbf{1}_{(a,b)}(x) dx}{2\pi x(x(1-\theta) + \theta(\alpha\theta - \alpha + 1))} + d_{\alpha, \theta} \mathbf{1}_{|\theta-1| > \frac{1}{\sqrt{\alpha}}} \delta_{x_{\alpha, \theta}}(dx),$$

where $a, b = (1 \mp \sqrt{\alpha})^2$ and where $x_{\alpha, \theta}$ and $d_{\alpha, \theta}$ are explicit constants.

ADDITIVE MODEL: OUTLIERS

When $\theta > 1$, we recover the following classical convergence:

- $\lambda_1(W_n) \xrightarrow{\mathbb{P}} \theta + 1/\theta$;
- $|\langle \phi_1(W_n), e_1 \rangle| \xrightarrow{\mathbb{P}} \sqrt{1 - 1/\theta^2}$.

MULTIPLICATIVE MODEL: OUTLIERS

When $\theta > 1 + 1/\sqrt{\alpha}$, we recover the following classical convergence:

- $\lambda_1(S_n) \xrightarrow{\mathbb{P}} x_{\alpha, \theta}$;
- $|\langle \phi_1(S_n), e_1 \rangle| \xrightarrow{\mathbb{P}} \sqrt{d_{\alpha, \theta}}$.

ADDITIVE MODEL: OVERLAPS

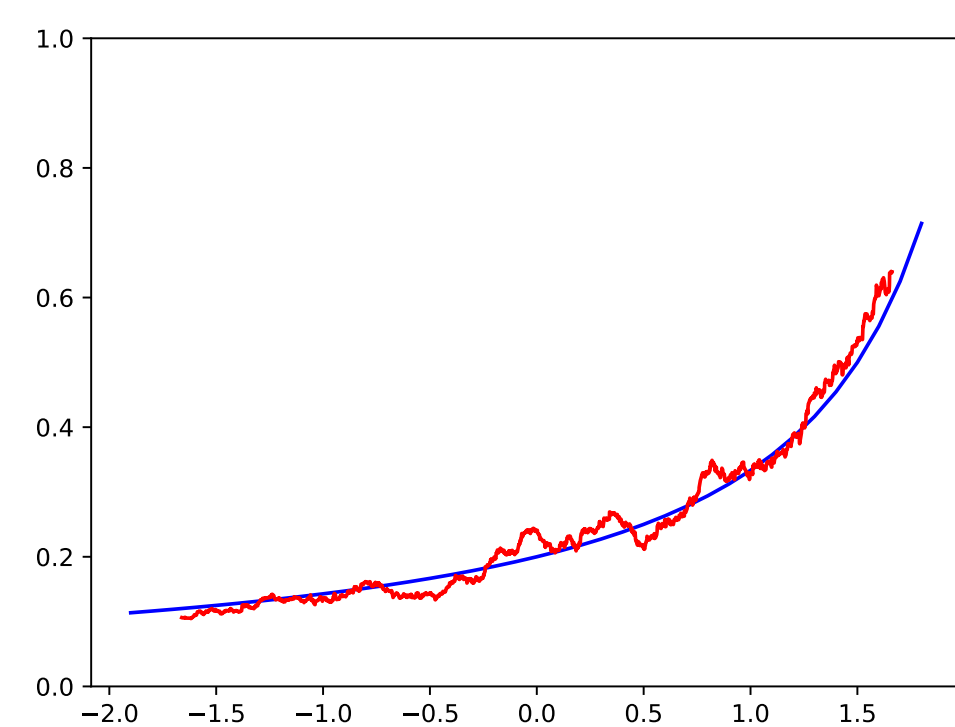
Let $x \in (-2, 2)$ and $\varepsilon > 0$. Then for any sequence $1 \gg \varepsilon_n \gg n^{-1/2+\varepsilon}$, denoting

$$\mathcal{I}_{\varepsilon_n}(x) = \{i : |\lambda_i(W_n) - x| \leq \varepsilon_n\},$$

it holds that

$$\frac{n}{|\mathcal{I}_{\varepsilon_n}(x)|} \sum_{i \in \mathcal{I}_{\varepsilon_n}(x)} |\langle \phi_i(W_n), e_1 \rangle|^2 \xrightarrow{\mathbb{P}} \frac{1}{\theta^2 - \theta x + 1}. \quad (1)$$

Simulations for $n = 4000$, $\theta = 2$:



MULTIPLICATIVE MODEL: OVERLAPS

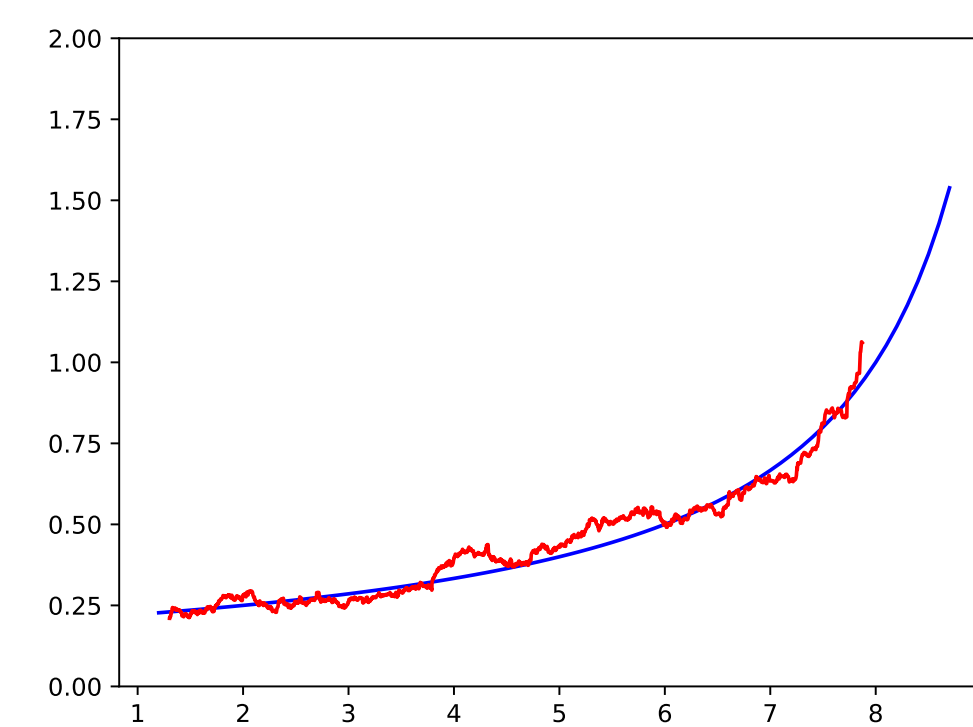
Let $x \in (a, b)$ and $\varepsilon > 0$. Then, for any sequence $1 \gg \varepsilon_n \gg n^{-1/2+\varepsilon}$, denoting

$$\mathcal{I}_{\varepsilon_n}(x) = \{i : |\lambda_i(S_n) - x| \leq \varepsilon_n\},$$

it holds that

$$\frac{n}{|\mathcal{I}_{\varepsilon_n}(x)|} \sum_{i \in \mathcal{I}_{\varepsilon_n}(x)} |\langle \phi_i(S_n), e_1 \rangle|^2 \xrightarrow{\mathbb{P}} \frac{\theta}{x(1-\theta) + \theta(\alpha\theta - \alpha + 1)}. \quad (2)$$

Simulations for $n = 2000$, $\alpha = 4$, $\theta = 2$:



SKETCH OF PROOFS

We focus on the Wigner case as the arguments are the same in the Wishart setting.

- **Outliers.** The largest eigenvalue $\lambda_1(n^{-1/2} X_n)$ converges in probability towards 2. When $\theta > 0$, the eigenvalues of W_n interlace with those of $n^{-1/2} X_n$. Therefore, only $\lambda_1(W_n)$ may be asymptotically larger than 2. It implies that the atom of $\mu_{sc, \theta}$ at $\theta + 1/\theta$, which exists whenever $\theta > 1$, is the limit of $\lambda_1(W_n)$ in probability. Moreover, its mass is the limit of the projection of $\phi_1(W_n)$ onto the direction of the spike.
- **Overlaps.** Let $x \in (-2, 2)$, define $\mathcal{I}_{\varepsilon_n}(x) := [x - \varepsilon_n, x + \varepsilon_n]$ and denote $f_{sc, \theta}$ the density of $\mu_{sc, \theta}$. The idea is to estimate $\mu_{(W_n, e_1)}(\mathcal{I}_{\varepsilon_n}(x))$ in two different ways. Firstly, $\mu_{(W_n, e_1)}(\mathcal{I}_{\varepsilon_n}(x)) \approx 2\varepsilon_n f_{sc, \theta}(x) + o_1(1)$ for a small error $o_1(1)$. Secondly, for a small error $o_2(1)$, if μ_{W_n} denotes the empirical spectral measure of W_n :

$$\mu_{(W_n, e_1)}(\mathcal{I}_{\varepsilon_n}(x)) = \left(\frac{n}{|\mathcal{I}_{\varepsilon_n}(x)|} \sum_{i \in \mathcal{I}_{\varepsilon_n}(x)} |\langle \phi_i(W_n), e_1 \rangle|^2 \right) \mu_{W_n}(\mathcal{I}_{\varepsilon_n}(x)) \approx \left(\frac{n}{|\mathcal{I}_{\varepsilon_n}(x)|} \sum_{i \in \mathcal{I}_{\varepsilon_n}(x)} |\langle \phi_i(W_n), e_1 \rangle|^2 \right) \times (2\varepsilon_n f_{sc, \theta}(x) + o_2(1)).$$

Using local laws obtained by Knowles and Yin [1], one can prove that $o_1(1)$ and $o_2(1)$ are of smaller order than ε_n with high probability. Therefore, the left-hand side of (1) converges in probability towards the ratio of the densities $f_{sc, \theta}(x)/f_{sc, 0}(x) = (\theta^2 - \theta x + 1)^{-1}$.

RELATED RESULTS

- **Generalizations.** The present results are taken from [4] and are particular instances of the study of the more general deformed models $W_n = n^{-1/2} X_n + A_n$ and $S_n = n^{-1} \Sigma_n^{1/2} X_n X_n^T \Sigma_n^{1/2}$, for general perturbations A_n and Σ_n whose normalized spectra respectively converge towards deterministic probability measures μ_A and μ_Σ . In these cases, the empirical spectral measures of W_n and S_n respectively converge towards the free convolution $\mu_{sc} \boxplus \mu_A$ and the free product $\mu_\alpha \boxtimes \mu_\Sigma$, where μ_{sc} is the semicircle law and μ_α the Marchenko-Pastur law with parameter α . If θ is an eigenvalue of A_n (resp. S_n) with associated eigenvector v_1 , the spectral measures in direction v_1 still converge towards deterministic probability measures $\mu_{sc, A, \theta}$ and $\mu_{\alpha, \Sigma, \theta}$, whose Stieltjes transforms are explicit functions of the Stieltjes transforms of $\mu_{sc} \boxplus \mu_A$ and $\mu_\alpha \boxtimes \mu_\Sigma$. In these settings, the outliers of W_n and S_n still correspond to the atoms of the limiting spectral measures and we prove the convergence of overlaps between the eigenvectors of W_n (resp. S_n) and v_1 , averaged on small scales, towards the ratio of the densities of $\mu_{sc, A, \theta}$ (resp. $\mu_{\alpha, \Sigma, \theta}$) and $\mu_{sc} \boxplus \mu_A$ (resp. $\mu_\alpha \boxtimes \mu_\Sigma$).
- **Other works.** Consider the general deformed models $W_n = n^{-1/2} X_n + A_n$ and $S_n = n^{-1} \Sigma_n^{1/2} X_n X_n^T \Sigma_n^{1/2}$ introduced above and suppose that θ is an eigenvalue of A_n (resp. S_n) with associated eigenvector v_1 . When θ belongs to the support of the asymptotic spectrum of A_n (resp. S_n), the convergence of the overlaps between the eigenvectors of W_n (resp. S_n) and v_1 is a *microscopic* confirmation of the results of Allez and Bouchaud [2] (in the Wigner setting) and Ledoit and P  ch   [3] (in the Wishart setting), who derived the asymptotic behavior of the overlaps $|\langle \phi_\lambda, v_\gamma \rangle|^2$ by taking the average over eigenvectors ϕ_λ associated to eigenvalues of W_n (resp. S_n) belonging to a *macroscopic* part of $\text{Supp}(\mu_{sc} \boxplus \mu_A)$ (resp. $\text{Supp}(\mu_\alpha \boxtimes \mu_\Sigma)$) and the average over eigenvectors v_γ of A_n (resp. S_n) belonging to a *macroscopic* part of the asymptotic spectrum of A_n (resp. S_n). When θ does not belong to the support of the asymptotic spectrum of A_n (resp. S_n), such a macroscopic average is not available whereas the spectral measure approach still works.

References

- [1] Antti Knowles and Jun Yin. Anisotropic local laws for random matrices. *Probab. Theory Related Fields*, 169(1-2):257–352, 2017.
- [2] Romain Allez and Jean-Philippe Bouchaud. Eigenvector dynamics under free addition. *Random Matrices Theory Appl.*, 3(3):1450010, 17, 2014.
- [3] Olivier Ledoit and Sandrine P  ch  . Eigenvectors of some large sample covariance matrix ensembles. *Probab. Theory Related Fields*, 151(1-2):233–264, 2011.
- [4] Nathan Noiry. Spectral Measures of Spiked Random Matrices. Preprint, arXiv:1903.11731, 2019.